



Early View

Review

Safety data in randomised real-world evidence studies: Salford Lung Study learnings

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Safety data in randomised real-world evidence studies: Salford Lung Study

learnings

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Take-home message (for social media): Using results from the Salford Lung Study, we demonstrate that the understanding of safety data from randomised real-world evidence studies can be optimised by analysing both actual treatment received and randomisation group.

Short, plain language summary: Clinical trials often have strict rules about who can and can't take part. This limits the types of people included in the study, for example by age or if they have other diseases. Because of this, researchers do not always know exactly how a medicine will work in each

individual. A type of trial, called a randomised real-world evidence (RRWE) trial, addresses some of these issues and can be used to see how medicines work in different people - for instance, the RRWE described here was designed so that the people involved were allowed to change their medication treatment throughout. Using this study, we looked at results to see how safe different medicines for asthma or COPD were for different people. We show that by looking at both the treatment assigned to people at the start of the trial, and at the medicines they actually took during the trial, researchers can make sure they are as accurate as possible when trying to find out how safe a medicine is.

ABSTRACT

Evidence to support clinical decision making must be based on safety data that have been captured, analysed, and interpreted in a robust and reliable way. Randomised real-world evidence (RRWE) studies provide the opportunity to evaluate the use of medicines in patients and settings representative of routine clinical practice. However, elements that underpin the design of RRWE studies can have a significant impact upon the analysis, interpretation, and implications of safety data.

In this narrative review, we use data from the Salford Lung Study; two prospective, 12-month, open-label, parallel-group, phase III randomised controlled trials conducted in primary care in the UK; to highlight the importance of capturing treatment modifications when attempting to evaluate safety events according to actual treatment exposure.

We demonstrate that analysing safety data by actual treatment received (i.e. accounting for the treatment modifications that occur routinely in the primary care setting) provides additional insight beyond analysing according to randomised treatment strategy only.

It is therefore proposed that understanding of safety data from RRWE trials can be optimised by analysing both by randomised group and by actual treatment received.

Introduction

Treatment guidelines and management algorithms for a variety of conditions, including the respiratory diseases chronic obstructive pulmonary disease (COPD) and asthma, are largely based on data from efficacy randomised controlled trials (RCTs) [1, 2]. Traditional double-blind RCTs typically recruit patients according to strict inclusion/exclusion criteria and are conducted under highly controlled conditions - so as to minimise bias due to confounding variables. However, such trials do not properly reflect how medicines are used in everyday clinical practice and under-represent characteristics within the wider disease population. Consequently, generalising findings from traditional double-blind RCTs to patients in routine clinical practice can be problematic [3, 4]. For example, a previous study reported that only 5.4% of asthma patients satisfied four commonly used criteria for entry to an asthma efficacy RCT (absence of co-morbidity, forced expiratory volume 50–85% of predicted, present or historical reversibility of 12% in the last year, non-smoker or, if ex-smoker, a smoke burden less than 10 pack years); the proportion of eligible patients further reduced to 3.3% when eligibility was made more stringent by the inclusion of two additional common criteria [3].

There is a need for evidence derived from studies conducted in conditions closer to routine clinical practice [4, 5]. RCTs conducted in primary care with elements of real-world design, referred to hereafter as randomised real-world evidence (RRWE) studies, provide much needed insights into the effectiveness and safety of medicines in patients and clinical settings that represent routine clinical practice. However, the apparent simplicity in the design of RRWE studies, paradoxically, adds complexity to the understanding of the safety data collected (outlined in Box 1). In this narrative review article, we discuss the analysis, interpretation and implications of safety data obtained in RRWE studies, using the phase III Salford Lung Study (SLS) as an illustrative example [6–10].

Safety data in the SLS

The SLS programme comprised two prospective, 12-month, open-label, parallel-group, phase III effectiveness RCTs conducted in primary care in and around Salford, United Kingdom (UK). The SLS enrolled a large population of patients with a general practitioner's (GP's) diagnosis of either COPD (SLS COPD) or asthma (SLS asthma). The trials evaluated the effectiveness and safety of initiating once-daily inhaled fluticasone furoate/vilanterol (FF/VI) compared with continuation of usual care (UC) (Figure 1) [9, 10]. Randomisation to initiate FF/VI or continue UC was stratified in both trials; specifically, patients were stratified by baseline maintenance therapy, to minimise confounding that might result from differences between the treatment groups.

In the SLS, more than 7000 patients provided written consent for their data to be used for safety monitoring [11]. Safety data (serious adverse events [SAEs], both treatment-related and non-treatment-related, and non-serious adverse drug reactions [ADRs; non-serious AEs considered to be treatment-related in the opinion of the investigator]) were collected by near real-time review of patients' electronic health records (EHRs) following electronic trigger alerts and monitoring of primary care data. Safety was assessed daily during the 4.5 years of the study. A detailed description of the safety monitoring and unique methods of data collection in the SLS has been published previously [12]. Crucially, as safety data could be monitored remotely using the EHR, study-specific visits were not mandated and patients could continue with their lives as usual - unlike in traditional efficacy RCTs where potentially inconvenient research site visits are required.

Safety endpoints included SAE frequency and type and there was specific examination of SAEs of special interest (SAESI), defined as known class effects of inhaled corticosteroids (ICS) and/or long-acting beta₂-agonists (LABA). SAESI groupings were defined *a priori* using groups of related Medical Dictionary for Regulatory Activities (MedDRA) event terms (by MedDRA structure, where possible), ensuring all relevant events were captured. One example was pneumonia, a known and accepted class effect of ICS use in patients with COPD. The pneumonia SAESI group was defined *a priori* by 69 MedDRA preferred terms; any verbatim term for an SAE that mapped to one of these terms was

therefore counted as a pneumonia SAESI. This ensured capture of all pneumonia events (*e.g.* those where an infective agent was part of the reported pneumonia term).

Understanding the impact of RRWE study design on safety data

Patient population

The heterogeneity of the patient population enrolled in RRWE studies has implications for understanding safety data. Including patients with multiple comorbidities and with many comorbid disease-related events can lead to far more SAEs than are usually seen in traditional RCTs, which typically exclude such patients. For example, the SLS did not exclude patients with a comorbid cancer diagnosis unless they were predicted to live for less than a year following randomisation. This resulted in more cancer-associated SAEs than would be observed in a traditional RCT. Although bias is possible in an open-label trial, randomisation reduces the risk for bias, as well as ensuring that reporting of SAEs which are non-treatment related should be balanced across groups, and indeed in the SLS there were no observed differences between treatment groups. Therefore, when SAEs are analysed, patient characteristics such as comorbidities should be considered.

Treatment modifications – analysing by actual versus randomised treatment

Modification of treatment is common in everyday clinical practice. This can be “stepping up” treatment to provide better disease control and management (*e.g.* addition of a new class of drug), or “stepping down” treatment in situations where disease control has been maintained for a specific period of time (*e.g.* removal of one class of drug) [13]. Modification could also involve switching to a different treatment altogether. Physicians and patients regularly modify treatment for asthma and COPD for various reasons, including symptom control, patient preference, familiarity, cost, reimbursement status, adverse reactions, convenience, device characteristics, adequate inhaler technique, disease progression and potential drug interactions. As such, there is a need to both allow and capture these treatment modifications in studies intended to more closely reflect usual

clinical practice. Accordingly, treatment modifications are an integral element in the design, execution and analytical approach of RRWE studies. By capturing treatment modifications, RRWE studies can provide evidence founded on actual treatment received (true exposure risk) [10], an approach that could be increasingly adopted in future effectiveness studies. Traditional RCTs generally do not allow for treatment modifications; participants must adhere to their randomised treatment throughout the study, and this is strictly monitored. Therefore, understanding the design of RRWE studies and whether participants have the option to modify their treatment directly impacts upon our ability to understand and interpret safety data.

Patients were able to modify their treatment in the SLS. This had far-reaching implications for the analysis and interpretation of the collected safety data. Patients who were randomised to the FF/VI group initiated treatment with FF/VI but were allowed to continue FF/VI or to change to any other suitable medicine for their condition while remaining in the study and being retained for analysis purposes in their original FF/VI randomisation group. Although randomised to initiate FF/VI, patients could have been taking a different medicine at any subsequent time (ICS, LABA, or long-acting muscarinic antagonist [LAMA]; either alone or in different combinations, according to the study protocol). If patients who modified their treatment continued with the new treatment for several months, it is questionable whether any later AEs can be considered part of the FF/VI “treatment period” rather than the “actual” treatment being taken at the time of the event. Here, the FF/VI randomisation group is not equivalent to an FF/VI treatment group; comparison of safety data between the randomisation groups in a RRWE study does not equate to a treatment comparison, but to a strategy comparison.

Analysis of safety data by actual treatment received is itself not without potential bias. Firstly, there could be a carry-over effect of risk; in the SLS for example, the carry-over of pneumonia risk, which was addressed in the actual treatment analysis by applying a risk window. Secondly, bias could be introduced as a result of the implications associated with asymmetric treatment modification. In the SLS, patients not content with FF/VI treatment will have switched to an

alternative treatment, resulting in a comparison of safety events between patients who remained on FF/VI throughout the study and a mixed group of patients. This latter group included patients who changed from FF/VI to an alternative treatment during the study, patients who continued the treatment they were receiving before the study, and also patients in the UC group who modified their prior study treatment in some way. The asymmetry of treatment modification in the SLS could favour treatment stability in the FF/VI group, but there is also an inherent stability for patients who remained on their prior (UC) medication during the study. Similarly, when analysing safety data by randomised treatment group, there is an inherent stability built into the UC comparator group *versus* the FF/VI randomisation group; the bias here will be for the stability in the UC group. Importantly, additional bias to the safety analysis could be introduced by treatment change temporally related to AEs or SAEs.

Use of a comparator group in RRWE studies – what constitutes usual care?

To ensure relevance to everyday clinical practice, the comparator for an investigational medicine in a RRWE study is typically a UC group. This group can receive multiple different drugs, administered alone or in a variety of combinations. In the SLS, the random allocation was patients being randomised either to initiate FF/VI or to continue their previous treatment (UC). For asthma and COPD, UC can mean a variety of treatments differing by class (*e.g.* ICS, LABA, LAMA alone or in combination, according to the study protocol), dose, or brand. In the SLS, patients randomised to the UC group were allowed at any time and for any reason, to modify their treatment to any other suitable UC medication (but not to FF/VI), whereas patients randomised to initiate FF/VI could modify their treatment to UC. The potential for continuous variation in the treatments being taken by patients randomised to the FF/VI and UC groups complicates data analysis: it precludes direct comparison of randomisation groups from being equated with the safety of one specific treatment relative to another.

Open-label design

When an open-label study seeks to compare a new medicine with existing therapies, patients and physicians will be aware that they are receiving something new, and this brings the potential for reporting bias [14]; it might be more likely that new AEs will be assigned causally to a new medicine than to an existing older one. Patients and physicians become familiar with asthma and COPD medications used for the long-term management of these conditions, and consequently, patients could establish preference for a particular medication or inhaler type. These potential issues associated with investigator evaluation and attribution of the causality for a given AE apply to both RRWE studies and traditional efficacy RCTs. In this respect, future studies may benefit from a blinded AE reporting team for causality and severity.

The issues surrounding reporting bias in open-label studies is potentially exemplified in the SLS, which commenced prior to licensing of FF/VI in the UK. Although a number of blinded studies have shown an AE profile for FF/VI similar to other ICS/LABA therapies, the SLS assigned causality of AEs to FF/VI use (*i.e.* more on-treatment ADRs/treatment-related AEs with FF/VI than for UC) (online supplementary table S1) [15, 16]. Whilst this could be a finding of the RRWE that was not reported in RCTs, it could also be a result of reporting bias. Further RRWEs could help to provide additional data to clarify this disparity. While the comparative number of SAEs in the different treatment groups should be considered, the incidence of events that are likely to be related to these medicines (based on the likely class effects of ICS or LABA – those selected as SAESI in the SLS) was comparable between randomisation groups (table S2) [15, 16].

Safety data in RRWE studies: examples from the SLS

Collecting information relating to treatment modifications was an important feature of the SLS study design. From the perspective of the randomisation groups, treatment modifications in the SLS were asymmetric. Patients randomised to FF/VI could change treatment to any other suitable treatment for any reason (*i.e.* commence treatment considered to be UC), while patients randomised to UC

were free to modify their treatment to any other UC treatment (except FF/VI). Any patient who modified their treatment remained in their original randomisation group for analysis purposes, including those who changed from treatment with FF/VI to another treatment.

In SLS COPD [9], 24% of patients randomised to initiate FF/VI underwent at least one treatment modification (figure 2); the majority changed back to their previous medicines but nevertheless remained in the FF/VI group. In the UC randomisation group, 11% of patients modified their treatment, but this meant that they remained on a UC treatment while remaining in the UC randomisation group. In SLS asthma [10], similar rates of treatment modifications were observed: 22% of patients in the FF/VI randomisation group and 18% in the UC randomisation group modified their treatment regimen (figure 2); again, most of the patients in the FF/VI group changed back to the medicine they were taking at study entry. These findings result from the open-label design of the SLS and possibly reflect the apparent preference of some patients and/or physicians for familiar treatments. Our experiences from the SLS illustrate the importance of capturing modifications of treatment when attempting to evaluate safety according to actual treatment exposure. Comparing randomisation groups is not the same as comparing the safety of FF/VI *versus* the safety of UC. It is comparing the strategy of initiating treatment with FF/VI *versus* the strategy of continuing UC.

In the SLS, we analysed safety data according to both randomisation group and actual treatment received. The former approach evaluated the effectiveness and safety of the strategy of initiating FF/VI (randomisation to the FF/VI group) *versus* continuing previous UC (randomisation to the UC group). The latter approach determined the safety of treatment being taken at the time of onset of the event, regardless of randomisation group. Here there is a juxtaposition of exposure risk and the occurrence of an AE; if a patient is not currently exposed to the treatment and has an event, it cannot reasonably be assigned to having occurred while “on treatment” with that product. However, this is also inadequate as some AEs might be associated with cumulative recent exposure, rather than immediate exposure - pneumonia for example. In the SLS, an “at-risk” window was pre-planned for the evaluation of pneumonia events; the relevant risk window of treatment at the time

of the event for pneumonia was the treatment taken for the majority of the 28-day period prior to onset of the event. Such rules for assigning safety events to a risk period of actual treatment need to be determined *a priori* in RRWE studies, with differences in time of treatment initiation/randomisation and AE onset, and possible drug exposure risk levels being considered in each case.

At the request of regulatory authorities, the SAESI of pneumonia was statistically compared to determine non-inferiority by randomisation group rather than by actual treatment in the SLS. Although characterising the incidence of SAESI of pneumonia had not been an original objective of the SLS, the real-world design and use from the outset of a comprehensive EHR system to monitor safety (which allowed access to all collected healthcare information from the point of patient consent) enabled pneumonia events to be collected both retrospectively and prospectively while the study was ongoing. In SLS COPD, there was a small numerical difference in the number of patients experiencing at least one pneumonia SAESI: 94 (7%) patients in the FF/VI randomisation group *versus* 83 (6%) in the UC randomisation group (table 1); statistically, being randomised to initiate FF/VI was not different to being randomised to continue UC, with an incidence ratio of 1.1 (0.9–1.5) [9, 15]. The total number of pneumonia SAESIs among both groups was 211; patients randomised to the FF/VI group and to the UC group experienced 104 (52%) and 97 (48%) events, respectively[15].

Pneumonia events can also be expressed as a rate per 1000 patient-years. In the SLS, this means a comparison of the rate of events per 1000 years of the strategy of randomising to initiate FF/VI *versus* the strategy of continuing UC. For events related to exposure, this is a useful comparison of rates. As discussed earlier, comparing randomisation groups in the SLS is not the same as comparing one treatment with another. In separate analyses of SLS COPD not previously published, the incidence of pneumonia SAESI was evaluated allowing for permitted treatment modifications that occurred during the study, *i.e.* analysis by actual treatment (for pneumonia, during the risk window). Patients in the FF/VI randomisation group, at different times, could have been receiving FF/VI, or have switched to another treatment. A more appropriate comparison is the

rate of events per 1000 patient-years of the actual treatment received at the time of the event. When actual treatments are compared in this way, the rates of pneumonia per 1000 patient-years were as follows: FF/VI 61.2 *versus* other ICS/LABA 64.0, and FF/VI + LAMA 60.9 *versus* other ICS/LABA/LAMA 92.7 (table 1) [15].

In SLS asthma, there was also a small observed difference in the number of patients with pneumonia SAESI when comparing randomisation groups: 23 (1%; 24 events) for FF/VI and 16 (<1%; 18 events) (table 1) [10, 16]. However, analysis by actual treatment, taking into consideration the treatment modifications during the study, revealed the same number of pneumonia events in both treatments (21 each) [10]. When pneumonia events were analysed by actual treatment risk the event rate per 1000 patient-years was: FF/VI 10.7 *versus* ICS 5.9 and ICS/LABA 9.7 (table 1) [10, 16].

The one-way treatment modification in the SLS that permitted modifying treatment from FF/VI to other treatments, but not allowing those randomised to UC to receive FF/VI, distorted the pneumonia rates and represented a potential bias against FF/VI when pneumonia SAESI data were analysed according to randomisation group. This was shown to be the case in SLS COPD when pneumonia was analysed by actual treatment received.

A similar phenomenon can be seen when examining other SAESI in the SLS, *e.g.* cardiovascular (CV) events, which were evaluated using standardised MedDRA outputs. In SLS COPD, when analysed by randomised treatment, 108 (8%) patients in the FF/VI randomisation group experienced a CV event *versus* 107 (8%) patients in the UC randomisation group (table 1) [9]. However, when analysed by actual treatment, rates per 1000 patient-years for the frequency of CV events were: FF/VI 110.1 *versus* other ICS/LABA 167.6, and FF/VI + LAMA 92.8 *versus* other ICS/LABA/LAMA 137.5 [15]. Whilst prior history of CV disease was not a stratification criterion at randomisation, it is suggestive that patients actually taking FF/VI were at lower risk than those taking other ICS/LABAs. Similarly, in SLS asthma, when analysed by randomisation group, 46 (2%) patients in both the FF/VI and UC randomisation groups had a CV event, but rates per 1000 patient-years of actual treatment were: FF/VI 23.3 *versus* ICS 31.6, and ICS/LABA 28.6 (table 1) [16]. This was also observed for SAESIs

of lower respiratory tract infections excluding pneumonia SAEs and incidence of decreased bone mineral density and associated fractures SAESIs.

This is an important observation. When treatment modifications are allowed in RRWE studies, analysis by randomisation group does not provide a comparison of the safety of one treatment *versus* another, and so analysis must be performed by actual treatment at the time of an event in order to understand the relative risk of the treatments. Although the comparison of pneumonia and CV events according to actual treatment resulted in fewer events with FF/VI compared with other classes of ICS/LABA (with the exception of pneumonia in SLS asthma, where the converse was true), this finding would have been overlooked if the study safety data had been analysed by randomisation group only.

Implications of safety data derived from RRWE studies

As noted earlier, data from the SLS highlights the potential for misinterpretation of safety data derived from RRWE studies. This is most apparent when analyses are conducted solely by randomised treatment group, and do not necessarily reflect true drug exposure during a study in which, by design, treatment modifications occur routinely. However, caveats also exist for analyses conducted according to actual treatment received, such as ignoring potential carry-over effects for risk and the potential biases associated with asymmetric treatment modification that could favour one treatment over another.

Regulators might insist on data analysis by randomised treatment group (the more familiar approach used in traditional efficacy RCTs), but this could be misleading for a RRWE study. It was for this reason that safety data from the SLS were presented using both approaches – by randomised treatment group (strategy) and actual treatment (risk exposure). We hope that our observations will facilitate better understanding of RRWE studies as they increase in number.

New methods for evaluating benefit/risk are needed for RRWE studies in which treatment modifications are permitted and are often the norm. Furthermore, the importance of providing

transparent reasoning and clear explanations for all key study design choices is emphasised by the inclusion of these points in the recommendations for prospective, observational, comparative effectiveness studies published by ISPOR [17].

The interpretation of safety data should be considered in the earliest stages of the design of RRWE studies and before preparation of the analysis plan. This is essential to ensure appropriate collection of data and its subsequent use. Unless safety data are handled correctly, a hazardous drug could potentially be mislabelled as safe or, alternatively, a safe drug blighted. A point also worth noting is that safety data from RRWE studies are unlikely to be suitable for integration into other types of analyses (*e.g.* meta-analyses, systematic reviews) due to inherent differences in trial design and the patient populations being studied.

As is the case for all interventional clinical studies, the SLS required expedited safety reporting of unexpected serious adverse reactions to regulatory authorities. Safety data from RRWE studies are also used in cumulative regulatory reporting through periodic benefit/risk evaluation reports, and as part of safety signal evaluation (actual treatment required). Safety data from RRWE studies will also be important for developing the safety profile of a drug, which requires safety data to be interpreted by actual treatment rather than the treatment strategy. If considered prior to study design, safety data from RRWE studies will also have value for evaluating benefit/risk in the setting of normal clinical practice.

Concluding remarks

RRWE studies provide the opportunity to evaluate medicines in patients and settings representative of routine clinical practice. The very elements that underpin the design of these studies can also impact significantly on the analysis, interpretation and implications of safety data collected in such trials. It is important to consider early in the design phase how best to evaluate safety elements in RRWE studies. As we have illustrated using data from the SLS, analysing safety by actual treatment received (*i.e.* accounting for the treatment modifications that occur in everyday primary care

settings) is essential. While safety data analysis by randomisation group provides useful additional insights into treatment strategy, we propose that the use of both approaches could represent the ideal understanding of safety data from RRWE studies.

If healthcare professionals are to be fully informed about the safety of the medicines they are prescribing to patients and how these will perform in the everyday clinical setting, the evidence informing their decision-making must include safety data that has been captured, analysed, and interpreted in a robust and reliable manner in real-world settings.

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Author contributions

CH, Safety Development Leader (GlaxoSmithKline plc.), led the manuscript writing together with SLS Steering/Publications Committee members (**AW**, **JV**, **NDB**, **JPN**) and the SLS study team (**CC**, **LF**, **CW**, **HE**, **NM**, **GC**, **SC**, **LJ**, **JF**). The data reported, some of which have been published previously [9, 10], were derived from the SLS trial dataset. **CH**, **AW**, **JV**, **CC**, **LF**, and **JPN**: study conception/design, data analysis/interpretation, and manuscript writing/review. **NDB** and **SC**: study conception/design, data acquisition, data analysis/interpretation, and manuscript writing/review. **CW** and **JF**: study conception/design, data acquisition, and manuscript writing/review. **HE** and **NM**: data acquisition and manuscript writing/review. **GC**: data acquisition, data analysis/interpretation, and manuscript writing/review. **LJ**: data analysis/interpretation and manuscript writing/review. All authors approved the final version of the manuscript for submission. **CH** is the manuscript guarantor, and attests that all listed authors meet authorship criteria and that no other authors meeting the criteria have been omitted.

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Conflict of interests

All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coi_disclosure.pdf and declare the following conflicts of interest: **CH, CC, LF, GC, LJ** and **JF** are all currently employed by GlaxoSmithKline plc.; they disclose non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study, as well as personal fees and other support from GlaxoSmithKline plc. outside of the submitted work. **SC** was an employee of GlaxoSmithKline plc. at the time of the study and has now retired; they disclose non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study, as well as personal fees and other support from GlaxoSmithKline plc. outside of the submitted work. **AW** discloses non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study, as well as speaker fees and expenses from GlaxoSmithKline plc. and advisory expenses from Chiesi outside of the submitted work. He is also Chairman of, and shareholder in, Reacta Biotech. **JV** discloses personal fees, non-financial support, and other support from GlaxoSmithKline plc. during the conduct of the study, as well personal fees from AstraZeneca, Boehringer Ingelheim, Chiesi, GlaxoSmithKline plc., and Novartis, and grants from Boehringer Ingelheim outside of the submitted work. **NDB** discloses non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study, as well as grants and personal fees from Almirall, AstraZeneca, GlaxoSmithKline plc. and Novartis, and other support from Boehringer Ingelheim outside of the submitted work. **JPN, HE and NM** disclose non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study. **CW** discloses non-financial support and other support from GlaxoSmithKline plc. during the conduct of the study, as well as personal fees from GlaxoSmithKline plc. outside of the submitted work.

Data sharing statement

Anonymised individual participant data from this study plus the annotated case report form, protocol, reporting and analysis plan, dataset specifications, raw dataset, analysis-ready dataset and clinical study report are available for research proposals approved by an independent review committee. Proposals should be submitted to www.clinicalstudydatarequest.com. A data access agreement will be required.

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<BOX 1>

Box 1. Study design features affecting interpretation of safety data in RRWE studies <i>versus</i> traditional efficacy RCTs	
RRWE Studies	Traditional Efficacy RCTs
<ul style="list-style-type: none">• Evaluate how medicines perform in routine practice• Broad inclusion criteria; includes patients with other comorbidities• Requires expedited safety reporting to regulatory authorities• Randomisation to treatment strategy• Generally open-label• Treatment modifications* permitted, as occurs in routine practice• Safety data analysed by actual treatment• Individual clinicians are responsible for data collection with minimal review of patients	<ul style="list-style-type: none">• Evaluate medicines under ideal and highly controlled conditions• Strict inclusion criteria; highly selected patient population• Requires expedited safety reporting to regulatory authorities• Randomisation to treatment groups• Generally double-blind• Treatment modifications* not permitted• Safety data analysed according to randomised treatment group• Multiple trained professionals are responsible for data collection with frequent patient reviews

RCT, randomised controlled trial; RRWE, randomised real-world evidence.

*Treatment modifications include a change in medication dose, a change in medication dose frequency, or changing to a different medication altogether.

TABLE 1 Pneumonia and CV effects SAESI analysed by randomised treatment group and by actual treatment in SLS COPD and SLS asthma.

SLS COPD			
SAESI	Randomised to continue UC group N=1403	Randomised to initiate FF/VI group N=1396	Actual treatment received* N=2799
Pneumonia			
Patients with event, n (%)	83 (6)	94 (7)	
Events, n [†]	97	104	
Incidence ratio (95% CI)	1.1 (0.9–1.5)		
Rate per 1000 patient-years (Total time at risk on current class of treatment, years [‡] / number of events / number of patients with event)			ICS/LABA/LAMA 92.7 (928.1/86/77) FF/VI + LAMA 60.9 (738.5/45/42) FF/VI 61.2 (556.0/34/31) ICS/LABA 64.0 (452.8/29/23) LAMA 29.9 (167.0/5/5) ICS 14.1 (70.8/1/1) ICS/LAMA 0.0 (38.3/0/0) LABA 0.0 (29.0/0/0) LABA/LAMA 39.0 (25.6/1/1)
CV effects			
Patients with event, n (%)	107 (8)	108 (8)	
Events, n	181	154	
Rate per 1000 patient-years (Total time at risk on current class of treatment, years [‡] / number of events / number of patients with event)			ICS/LABA/LAMA 137.5 (872.6/120/74) FF/VI + LAMA 92.8 (679.1/63/48) FF/VI 110.1 (508.5/56/40) ICS/LABA 167.6 (423.7/71/37) LAMA 90.1 (155.4/14/11) ICS 15.2 (65.6/1/1) ICS/LAMA 84.2 (35.6/3/2) LABA 185.9 (26.9/5/4) LABA/LAMA 83.1 (24.1/2/2)
SLS asthma			
SAESI	Randomised to continue UC group N=2119	Randomised to initiate FF/VI group N=2114	Actual treatment received N=4233
Pneumonia			
Patients with event, n (%)	16 (<1)	23 (1)	
Events, n	18	24	

Incidence ratio (95% CI)	1.4 (0.8–2.7)		
Rate per 1000 patient-years (Total time at risk on current class of treatment, years [†] / number of events / number of patients with event)			FF/VI 10.7 (1967.8/21/20) ICS 5.9 (849.7/5/5) ICS/LABA 9.7 (1642.9/16/14) ICS or ICS/LABA 8.4 (2492.6/21/19) FF/VI + ICS 0.0 (0.1/0/0)
CV effects			
Patients with event, n (%)	46 (2)	46 (2)	
Events, n	56	56	
Rate per 1000 patient-years (Total time at risk on current class of treatment, years [†] / number of events / number of patients with event)			FF/VI 23.3 (1805.7/42/35) ICS 31.6 (791.4/25/21) ICS/LABA 28.6 ICS or ICS/LABA (1538.8/44/35) FF/VI + ICS 29.6 (2330.2/69/56) 0.0 (0.1/0/0)

*The dose of FF/VI was either 100 µg or 200 µg FF with 25 µg VI – the dose of other therapies was per optimised usual care.

†Some patients experienced >1 pneumonia SAESI

‡Rate given for analysis by actual treatment (*i.e.* rate per 1000 patient-years based on a 28-day window for pneumonia and 1-day window for CV).

CI, confidence interval; COPD, chronic obstructive pulmonary disease; CV, cardiovascular; FF/VI, fluticasone furoate/vilanterol; ICS, inhaled corticosteroid; LABA, long-acting beta₂-agonist; LAMA, long-acting muscarinic antagonist; SAESI, serious adverse event of special interest; SLS, Salford Lung Study; UC, usual care

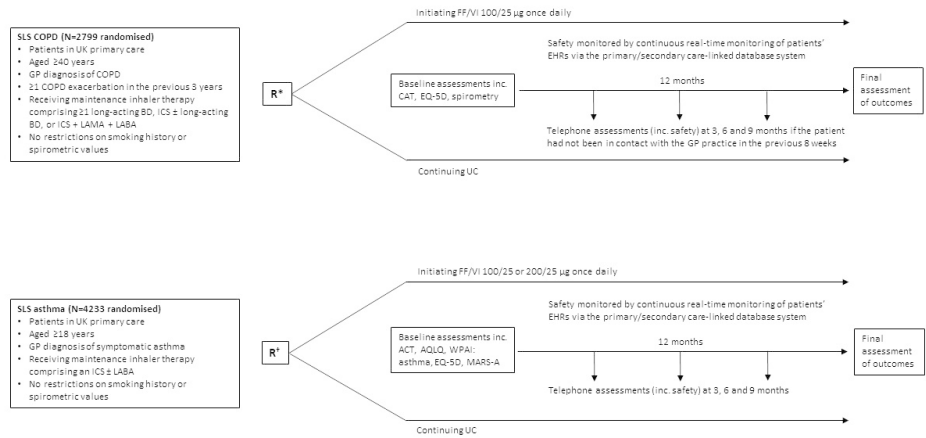


FIGURE 1 Overview of design of SLS COPD and SLS asthma.

*Randomisation stratified by presence/absence of a COPD exacerbation in the previous 12 months and baseline intended COPD maintenance therapy.

†Randomisation stratified by baseline ACT score (≥ 20 , 16–19, ≤ 15) and baseline intended asthma maintenance therapy.

ACT, Asthma Control Test; AQLQ, Asthma Quality of Life Questionnaire; BD, bronchodilator; CAT, COPD Assessment Test; COPD, chronic obstructive pulmonary disease; EHR, electronic health record; FF/VI, fluticasone furoate/vilanterol; GP, general practitioner; ICS, inhaled corticosteroid; LABA, long-acting beta2-agonist; LAMA, long-acting muscarinic antagonist; MARS-A, Medical Adherence Report Scale for Asthma; SLS, Salford Lung Study; UC, usual care; UK, United Kingdom; WPAI: asthma, Work Productivity and Activity Impairment questionnaire.

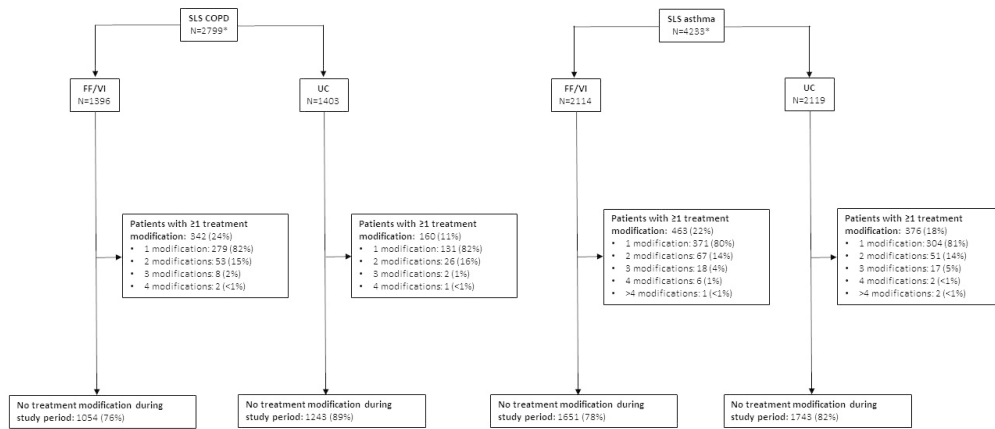


FIGURE 2 Treatment modifications by randomised treatment group during the 12-month study period in SLS COPD and SLS asthma.

A key design characteristic of the SLS was the ability for patients' treatment to be modified during the 12-month study period. Treatment modifications included a change in class of medication, an increase/decrease in medication dose, or a change in brand of medication. Patients who initiated treatment with FF/VI were permitted to modify treatment to UC; however, patients who initiated treatment with UC were not permitted to initiate treatment with FF/VI.

*Total study population (intent-to-treat).

COPD, chronic obstructive pulmonary disease; FF/VI, fluticasone furoate/vilanterol; SLS, Salford Lung Study; UC, usual care.

SUPPLEMENTARY MATERIAL

TABLE S1 On-treatment ADRs by system organ class by randomisation group in SLS COPD and SLS asthma ($\geq 1\%$ patients in either treatment group; ITT populations).

	SLS COPD		SLS asthma	
	UC	FF/VI	UC	FF/VI
System organ class, n (%)	N=1403	N=1396	N=2119	N=2114
Patients with any ADR	97 (7)	211 (15)	109 (5)	326 (15)
Infections and infestations	51 (4)	74 (5)	37 (2)	76 (4)
Respiratory, thoracic and mediastinal disorders	20 (1)	78 (6)	40 (2)	173 (8)
Gastrointestinal disorders	9 (<1)	22 (2)	15 (<1)	43 (2)
Nervous system disorders	3 (<1)	28 (2)	8 (<1)	28 (1)
Skin and subcutaneous tissue disorders	3 (<1)	14 (1)	6 (<1)	13 (<1)

ADR, adverse drug reaction; COPD, chronic obstructive pulmonary disease; FF/VI, fluticasone furoate/vilanterol; ITT, intent-to-treat; SLS, Salford Lung Study; UC, usual care.

TABLE S2 Summary of on-treatment serious ADRs of special interest by randomisation group in SLS COPD and SLS asthma (ITT populations).

	SLS COPD		SLS asthma	
	UC	FF/VI	UC	FF/VI
Special interest group, n (%)	N=1403	N=1396	N=2119	N=2114
Pneumonia	4 (<1)	4 (<1)	0	6 (<1)
Cardiovascular effects	3 (<1)	3 (<1)	2 (<1)	0
Cardiac arrhythmia	2 (<1)	3 (<1)	2 (<1)	0
Stroke	1 (<1)	0	0	0
Asthma/bronchospasm	0	0	0	2 (<1)
LTRI excluding pneumonia	1 (<1)	3 (<1)	0	1 (<1)
Effects on glucose	0	2 (<1)	0	2 (<1)
Effects on potassium	0	1 (<1)	0	0
Hypersensitivity	0	1 (<1)	0	0

ADR, adverse drug reaction; COPD, chronic obstructive pulmonary disease; FF/VI, fluticasone furoate/vilanterol; ITT, intent-to-treat; LRTI, lower respiratory tract infection; SLS, Salford Lung Study; UC, usual care.